

# Statistical analysis for the Johnson-Mehl germination-growth model

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- Studying the Johnson-Mehl germination-growth model in  $\mathbb{R}^d$ .
- Estimating parameters of specific parametric models for the conditional intensity by maximum likelihood.
- Model checking by new functional summary statistics related to the inhomogeneous K- function and to the Palm distribution of the typical Johnson-Mehl cell.



# Outline of Topics

- 1 Johnson-Mehl germination-growth Model
- 2 First and second-order properties
- 3 Functional summary statistics and non-parametric estimation
- 4 Parametric Models
- 5 Likelihood Analysis
- 6 A case study: Neurotransmitter data



# Definition

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- Time  $T((x, t), y)$

$$T((x, t), y) = t + \|x - y\|/v, \quad (x, t) \in \mathbb{R}^d \times [0, \infty) \text{ and } y \in \mathbb{R}^d.$$



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- New points and cells form and grow only in uncovered space.



## Definition(cont.)

- Growth ceases for each cell whenever and wherever it touches a neighbouring cell.



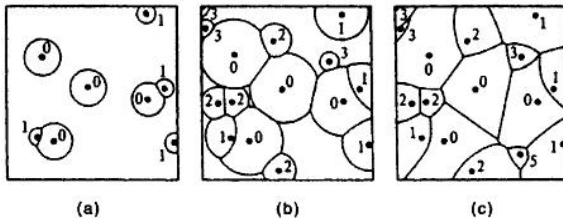
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- Cells:  $C_i = C(x_i, t_i) = \{y \in \mathbb{R}^d : T_i(y) \leq T_j(y) \text{ for all } j \neq i \text{ with } (x_j, t_j) \in \Psi\}$ , where  $T_j(x_i) = T((x_j, t_j), x_i)$ .



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**Figure:** The Johnson-Mehl model for times (a)  $t=1$ , (b)  $t=3$ , and (c)  $t=7$

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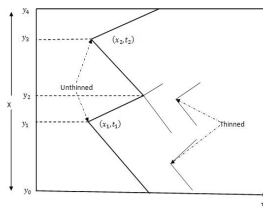


Figure: Thinned and unthinned points of a germination-growth process



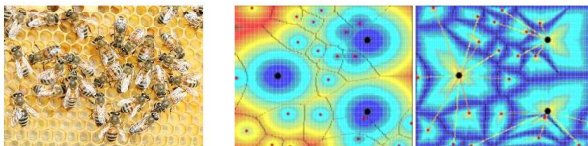
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**Figure:** Voronoi tessellation; cells are convex polyhedra.  
J-M tessellation: cells here are non-convex sets with curved boundaries

# Historical point of view

- J-M germination-growth model well studied from a probabilistic point of view, with the pioneering work by Kolmogorov (1937) and Johnson and Mehl (1939)



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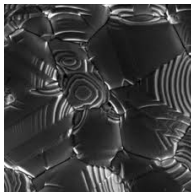


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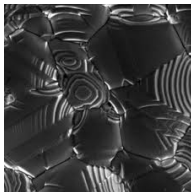


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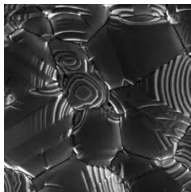


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- The statistical aspects?

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- The intensity of  $\Psi$  is given by

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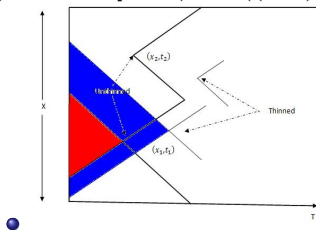


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**Figure:** Cone generated by  $(x_1, t_1)$  (Red colored area) and by the thinned point (Blue+Red colored area)

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- By (1) we have obtained a second-order non-linear differential equation with solution

$$\kappa(t) = \frac{c_1}{\cos^2(c_2t + c_3)}, \quad (2)$$

$c_1$ ,  $c_2$  and  $c_3$  are constants (and  $c_2t + c_3 \neq k\pi/2$ ,  $k \in \mathbb{Z} \setminus \{0\}$ )



## Second-order properties

- Let  $(x, s) \neq (y, t)$  in  $\mathbb{R}^d \times [0, \infty)$  with distance  $r = \|x - y\|$ ,  $(x, s)$  and  $(y, t)$  are in  $\Psi$ , if  $T((x, s), y) > t$  and  $T((y, t), x) > s$ , or  $r > v|s - t|$ .



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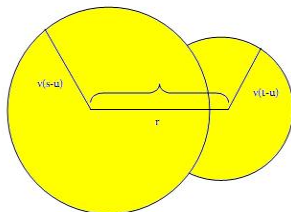


Figure:  $V_{\cup}(r, s - u, t - u)$



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- The pair correlation function is given by

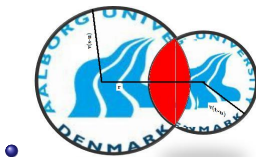
$$g(r, s, t) = \mathbf{1}[v|s - t| < r < v(s + t)] \exp \left( - \int_0^{(s+t-r/v)/2} \kappa(u) V_{\cap}(r, s - u, t - u) \, du \right) \\ + \mathbf{1}[r \geq v(s + t)].$$



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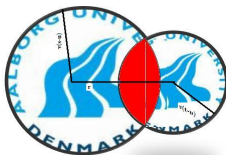
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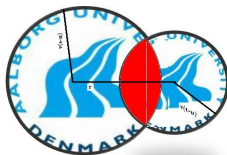


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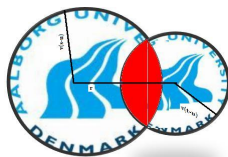
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- Since  $g(r, s, t)$  is not a function of  $r$  and  $s - t$  only. Therefore,  $\Psi$  is not second-order intensity-reweighted stationary.



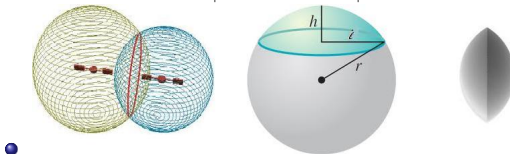
$$V_{\cap}(r, s - u, t - u)$$

- $V_{\cup}(r, s - u, t - u) = \omega_d[v(s - u)]_+^d + \omega_d[v(t - u)]_+^d - V_{\cap}(r, s - u, t - u)$



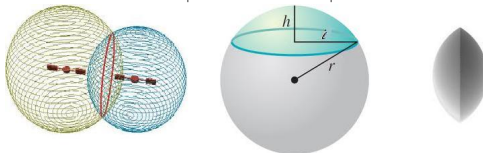
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- volume of a d-dimensional hyper-spherical cap:

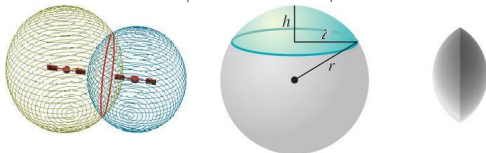
$$V_d(l, h) = \frac{1}{2} \frac{\pi^{d/2}}{\Gamma(1 + d/2)} r^d I_{(2lh - h^2)/l^2}((d + 1)/2, 1/2)$$

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- $$V_{\cap}(r, s - u, t - u) = V_d \left( v(s - u), \frac{[v(t - u)]^2 - (r - v(s - u))^2}{2r} \right) + V_d \left( v(t - u), \frac{[v(s - u)]^2 - (r - v(t - u))^2}{2r} \right).$$

# Inhomogeneous $K$ -function-like summary statistics and their Non-parametric Estimation

- For  $R > 0$ , define

$$K_1(R) = \mathbb{E} \sum_{i \neq j} \frac{\mathbf{1} \left[ x_i \in W, v(t_i + t_j) < \|x_i - x_j\| \leq R \right]}{|W| \rho(t_i) \rho(t_j)}$$



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- An unbiased estimator of  $K_1(R)$ :

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## Summary statistics based on the characteristics for the typical Johnson-Mehl cell

- The Palm distribution of the typical cell  $\mathcal{C}$  is defined by

$$\zeta|W|P(\mathcal{C} \in F) = \mathbb{E} \sum_i \mathbf{1} [T_j(x_i) > t_i \forall j \neq i, x_i \in W, C_i - x_i \in F]$$



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- Let  $o$  denote the origin.

$$C(o, t|\Phi) = \{y \in \mathbb{R}^d : T((o, t), y) \leq T((x_j, t_j), y) \text{ for all } (x_j, t_j) \in \Psi\}$$





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- Hence, by Slivnyak-Mecke formula

$$P(\mathcal{C} \in F) = \int P(T_j(o) > t \forall j, C((o, t)|\Phi) \in F) \kappa(t) dt / \zeta.$$



Palm distribution of the typical shortest nucleus-boundary distance  $R$ 

- In the Johnson-Mehl case, the distribution function for  $R$  is

$$D(r) = P(R \leq r) = 1 - \iint P(\Phi \cap H(t, r) = \emptyset) \kappa(t) dx dt / \zeta, \quad r > 0,$$



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$$D(r) = P(R \leq r) = 1 - \iint P(\Phi \cap H(t, r) = \emptyset) \kappa(t) dx dt / \zeta, \quad r > 0,$$

- 

$$H(t, r) = \{(y, u) \in \mathbb{R}^d \times [0, t] : \|y\| \leq 2r + v(t - u)\} \\ \cup \{(y, u) \in \mathbb{R}^d \times (t, t + r/v) : v(t - u) \leq \|y\| \leq 2r + v(t - u)\}.$$

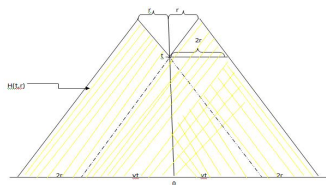


Figure: Example of the region  $H(t, r)$ .

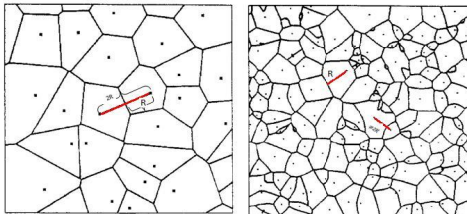
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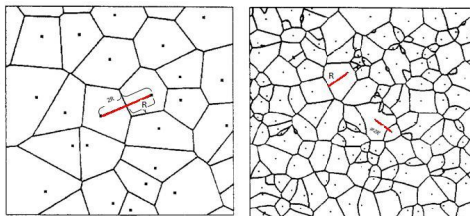
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**Figure:** Shortest boundary distance: Vorronoi tessellation (left panel), and Johnson-Mehl tessellation (right panel)

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**Figure:** Shortest boundary distance: Voronoi tessellation (left panel), and Johnson-Mehl tessellation (right panel)

- By ignoring edge effects, a ratio unbiased non-parametric estimate of  $D(r)$  is

$$\hat{D}(r) = \frac{1}{|W|} \sum_i \frac{\mathbf{1}[T_j(x_i) > t_i \forall j \neq i, x_i \in W, R_i \leq r]}{\hat{\zeta}}$$

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- From a statistical point of view in: only paper is Quine and Robinson (1992). They considered only the one-dimensional case  $d = 1$  and the time-homogeneous case,  $\beta = 1$ .



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- Chiu (1995) studied the limiting distribution of the time of completion for Johnson-Mehl model within a bounded region.



# Functional summary statistics for model M1

- The intensity function:  $\rho(t) = \exp(-\alpha\omega_d v^d t^{\beta+d} B(\beta, d+1)) \alpha t^{\beta-1}$





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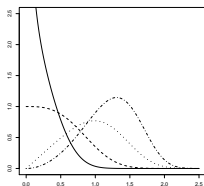
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**Figure:** Behavior of  $\rho(t)$  for model M1 with  $d = 2$  and  $\alpha = v = 1$ , when  $\beta = 0.5$  (solid line),  $\beta = 1$  (dashed line),  $\beta = 2$  (dotted line), and  $\beta = 3$  (dot-dashed line).

# Pair correlation function under model M1

- For  $d = 1$ ,  $V_{\cap}(r, s - u, t - u) = v(s + t - 2u) - r$ .
- The pair correlation function:

$$g(r, s, t) = 1[r > v|s - t|] \exp\left(\frac{\alpha(v(\beta + 1) + \beta)}{\beta(\beta + 1)2^{\beta}}(s + t - r/v)^{\beta+1}\right)$$



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$$P(R \leq r) = 1 - \int \exp\left(\frac{-2\alpha}{\beta} \left(2r(t + r/v)^{\beta} + \frac{v}{\beta + 1}t^{\beta+1}\right)\right) \alpha t^{\beta-1} dt / \zeta.$$



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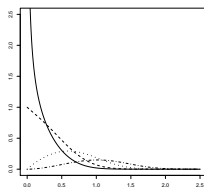
$$\rho(t) = \exp \left( -2\alpha v \left( t\Gamma(t; \beta, \gamma) - \frac{\beta}{\gamma} \Gamma(t; \beta + 1, \gamma) \right) \right) \frac{\alpha \gamma^\beta}{\Gamma(\beta)} t^{\beta-1} \exp(-\gamma t) \quad \text{if } d = 1$$



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**Figure:** Behavior of  $\rho(t)$  for model M2 with  $d = 2$ ,  $\alpha = v = \gamma = 1$ ,  $\beta = 0.5$  (solid line),  $\beta = 1$  (dashed line),  $\beta = 2$  (dotted line) and  $\beta = 3$  (dot-dashed line).



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Likelihood when  $\Phi$  is defined on  $W \times [0, \infty)$ 

- Assume  $d = 1$  and let  $\kappa = \kappa_\theta$  depends on a parameter  $\theta$ , e.g.,  $\theta = (\alpha, \beta) \in [0, \infty)^2$  in case of M1 or  $\theta = (\alpha, \beta, \gamma) \in [0, \infty)^3$  in case of M2.



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$$\lambda(x, t | \mathcal{H}_t) dt = \mathbf{1}[T_i(x) > t \forall t_i < t \text{ with } (x_i, t_i) \in \Psi_W] \mathcal{K}(dt), \quad (x, t) \in W \times [0, \infty),$$

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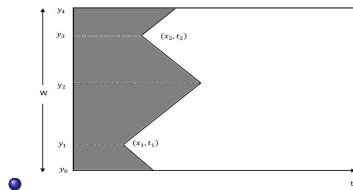


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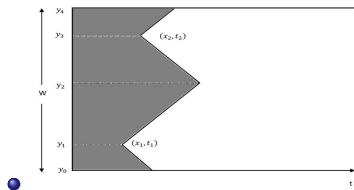


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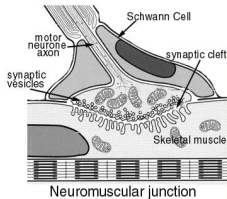
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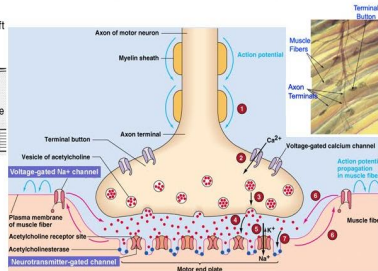


## Release of neurotransmitter at the neuromuscular junction

- The neuronal axon terminal at the neuromuscular junction has branches consisting of strands containing many randomly scattered sites.

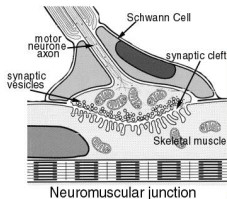


### The Neuromuscular Junction

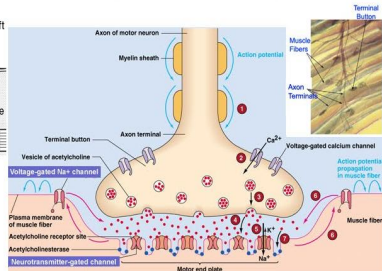


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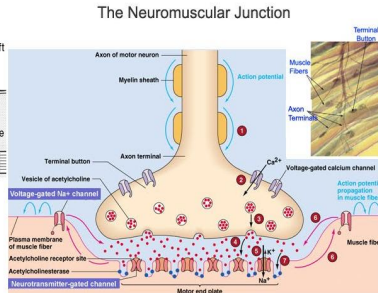
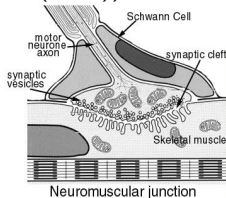


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- Each quantum released is assumed to cause release of an **inhibitory substance** which diffuses along the terminal at a **constant rate preventing further releases in the inhibited region** (Bennett and Robinson (1990)).



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- Among the transformed data, four outliers above 1 are deleted.
- Finally, 746 experiments with 101 experiments with no germinated seed and 645 with at least one germinated seed are obtained. The frequencies of 1's, ..., 4's now being 387, 210, 45, 3, respectively.



## Model checking

- Estimates:  $\hat{\alpha} = 1.29$ ,  $\hat{\gamma} = 13.3$ ,  $\hat{\beta} = 5.36$ ,  $\hat{v} = 0.018$ .



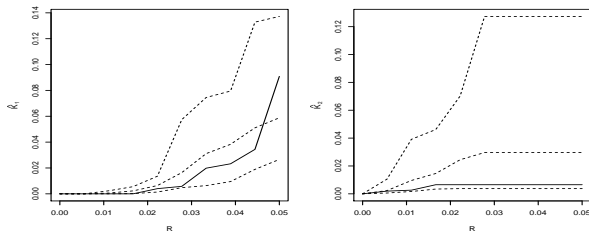
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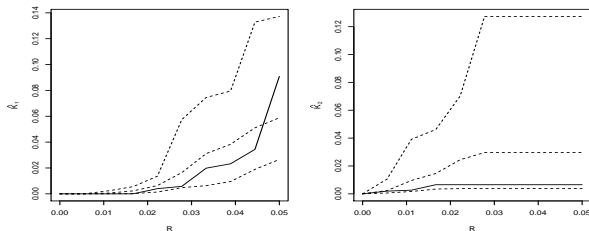
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**Figure:** Left: Estimated  $K_1$ -function for the data (solid line), and average and envelopes calculated from 39 simulations of the fitted model (dashed lines). Right: as left for  $K_2$ -function.

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**Figure:** Left: Estimated  $K_1$ -function for the data (solid line), and average and envelopes calculated from 39 simulations of the fitted model (dashed lines). Right: as left for  $K_2$ -function.

- For all  $R$  values,  $\hat{K}_1$  and  $\hat{K}_2$  for the data is between the envelopes, so the plot is in favor of the fitted model.

# Work in progress

- Estimating the parameters of model M1 by MLE



# Work in progress

- Estimating the parameters of model M1 by MLE
- Checking the fitted model by  $D(r)$





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