

Modeling group dispersal of particles with a spatiotemporal point process

Samuel Soubeyrand
INRA – Biostatistics and Spatial Processes
Joint work with L. Roques, J. Coville and J. Fayard

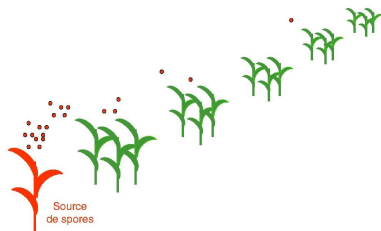


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Group Dispersal Project – Plant Health and Environment Dpt.

Spatiotemporal point processes in propagation models

Object of interest: species spreading using small particles (spores, pollens, seeds...)



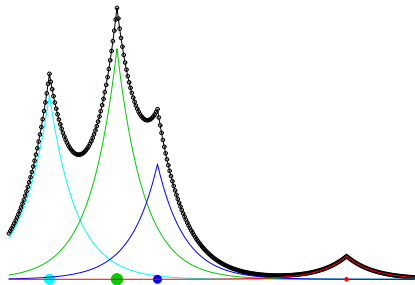
Sources of particles generate a spatially structured *rain* of particles

- ▶ rain of particles → spatial point process
- ▶ spatial structure → inhomogeneous intensity of the process

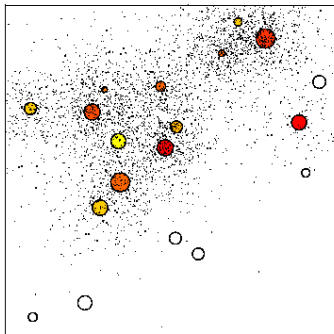
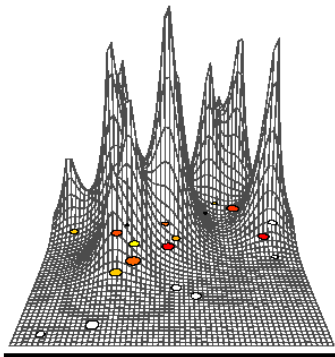
Intensity of the spatial point process formed by the deposit locations of the particles

The intensity is a convolution between

- ▶ the source process (spatial pattern and strengths) and
- ▶ a parametric dispersal kernel



Simulation of an epidemics



Dispersal kernel

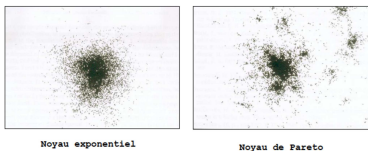
Dispersal kernel: probability density function of the deposit location of a particle released at the origin

The shape of the kernel is a major topic in dispersal studies: it determines

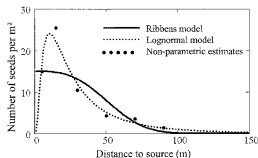
- ▶ the propagation speed
- ▶ the spatial structure of the population
- ▶ the genetic structure of the population

Main characteristics of dispersal kernels:

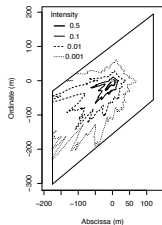
- ▶ long distance dispersal (Minogue, 1989)



- ▶ non-monotonicity (Stoyan and Wagner, 2001)

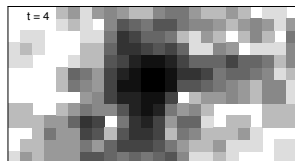
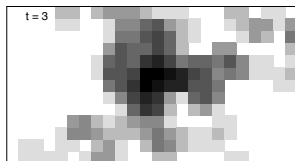
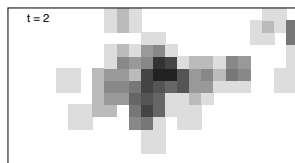
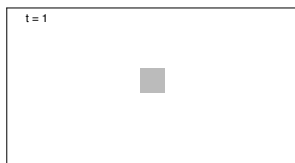


- ▶ anisotropy



Observation of secondary foci (clusters) in real epidemics

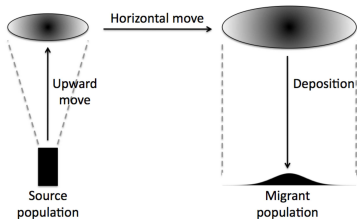
Epidemics of yellow rust of wheat in an experimental field (I. Satche)



- ▶ Classical justifications for patterns with multiple foci:
 - ▶ long distance dispersal
 - ▶ spatial heterogeneity
 - ▶ super-spreaders (a few individuals which infects many susceptible individuals)

- ▶ Classical justifications for patterns with multiple foci:
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- ▶ An other justification to be investigated: **Group dispersal**
 - ▶ Groups of particles are released due to wind gusts
 - ▶ Particles of any group are transported in an expanding air volume
 - ▶ At a given stopping time, particles of any group are projected to the ground



Group Dispersal Model (GDM): Spatial case

Deposit equation for particles:

A single point source of particles located at the origin of \mathbb{R}^2

J : number of groups of particles released by the source

N_j : number of particles in group $j \in \{1, \dots, J\}$

X_{jn} : deposit location of the n th particle of group j satisfying

$$X_{jn} = X_j + B_{jn}(\nu \|X_j\|), \quad (1)$$

where

X_j : final location of the center of group j ,

B_{jn} : Brownian motion describing the relative movement of the n th particle in group j with respect to the group center

ν : positive parameter

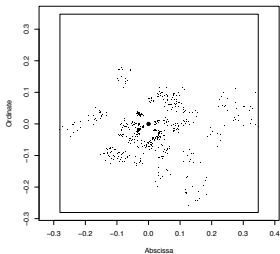
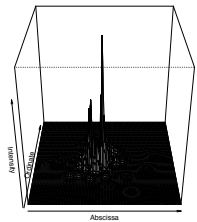
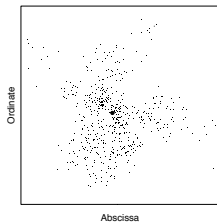
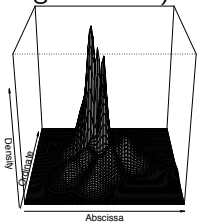
Assumptions about the deposit equation

- ▶ The random variables J , N_j , X_j and the random processes $\{B_{jn} : n = 1, \dots, N_j\}$ are mutually independent
- ▶ Number of groups: $J \sim \text{Poisson}(\lambda)$
- ▶ Number of particles in group j : $N_j \sim_{\text{indep}} P_{\mu, \sigma^2}(\cdot)$
- ▶ Group center location: $X_j \sim_{\text{indep}} f_{X_j}(\cdot)$
(features of f_{X_j} : decrease at the origin is more or less steep, tail more or less heavy, shape more or less anisotropic...)
- ▶ The Brownian motions B_{jn} are centered, independent and with independent components
They are stopped at time $t = \nu \|X_j\|$.
Then,

$$B_{jn}(\nu \|X_j\|) \sim_{\text{indep}} N(0, \nu \|X_j\| I)$$

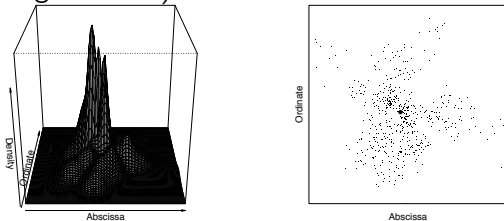
Dispersal from a single source

- ▶ Simulations: (Interpretation: Cox process or Neyman-Scott with double nonstationarity — in the center pattern and the offspring diffusion)



Dispersal from a single source

- ▶ Simulations: (Interpretation: Cox process or Neyman-Scott with double nonstationarity — in the center pattern and the offspring diffusion)



- ▶ Marginal probability density function (dispersal kernel):

$$f_{X_{jn}}(x) = \int_{\mathbb{R}^2} f_{X_{jn}|X_j}(x | y) f_{X_j}(y) dy = \int_{\mathbb{R}^2} \phi_{\nu,y}(x) f_{X_j}(y) dy.$$

The particles are n.i.i.d. from this p.d.f. while in the classical dispersal models the particles are i.i.d. from a dispersal kernel which may be of the form of f_{X_j} or $f_{X_{jn}}$

Discrepancies from independent dispersal

The GDM is compared with two independent dispersal models (IDM)

- ▶ IDM1: the number of particles in each group is assumed to be one. Thus, particles are independently drawn under the p.d.f. $f_{X_{jn}}$.
- ▶ IDM2: the number of particles in each group is assumed to be one and the Brownian motions are deleted (i.e. $\nu = 0$). Thus, particles are independently drawn under the p.d.f. f_{X_j} .

Moments

X : Deposit location of a particle

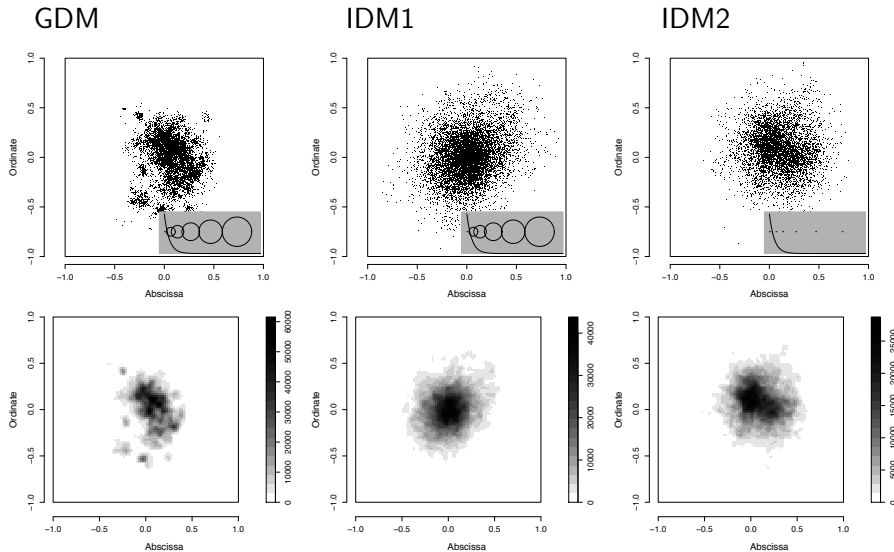
$Q(x + dx)$: Count of points in $x + dx$

Criterion	Model	Value
$E(X)$	GDM	$\begin{pmatrix} 0 \\ 0 \end{pmatrix}$
	IDM1	$\begin{pmatrix} 0 \\ 0 \end{pmatrix}$
	IDM2	$\begin{pmatrix} 0 \\ 0 \end{pmatrix}$
$V(X)$	GDM	$V(X_j) + \nu E(\ X_j\)l$
	IDM1	$V(X_j) + \nu E(\ X_j\)l$
	IDM2	$V(X_j)$
$E(\ X\ ^2)$	GDM	$E(\ X_j\ ^2) + 2\nu E(\ X_j\)$
	IDM1	$E(\ X_j\ ^2) + 2\nu E(\ X_j\)$
	IDM2	$E(\ X_j\ ^2)$
$E\{Q(x + dx)\}$	GDM	$\lambda \mu f_{X_{j_n}}(x) dx$
	IDM1	$\lambda f_{X_{j_n}}(x) dx$
	IDM2	$\lambda f_{X_j}(x) dx$
$V\{Q(x + dx)\}$	GDM	$\lambda[\mu f_{X_{j_n}}(x) dx + (\sigma^2 + \mu^2 - \mu) E\{\phi_{\nu, X_j}(x)^2\} (dx)^2]$
	IDM1	$\lambda f_{X_{j_n}}(x) dx$
	IDM2	$\lambda f_{X_j}(x) dx$
$\text{cov}\{Q(x_1 + dx), Q(x_2 + dx)\}$	GDM	$\lambda(\sigma^2 + \mu^2 - \mu) E\{\phi_{\nu, X_j}(x_1) \phi_{\nu, X_j}(x_2)\} (dx)^2$
	IDM1	0
	IDM2	0

GDM: larger variance of $Q(x + dx)$ and positive covariance
(decreasing with distance)
→ clusters (even with $\mu = 1$)

We expect multiple foci in the spatio-temporel case

Group dispersal model: Spatio-temporal case

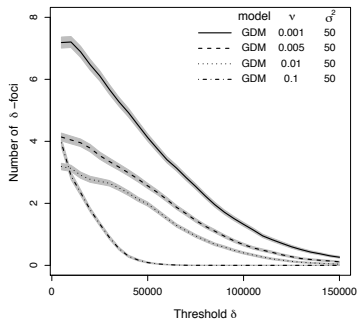
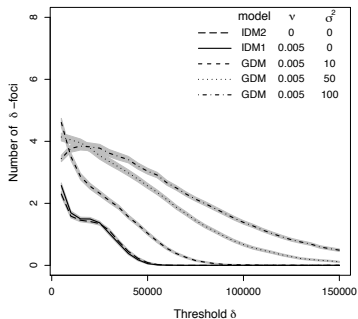


→ multiple foci under the GDM

Simulation study of the number of foci:

Definition

A δ -focus is a set of cells (from a regular grid) which are connected and whose intensity of points is larger than δ



Farthest particle (link with propagation speed)

Definition

The maximum dispersal distance during one generation is

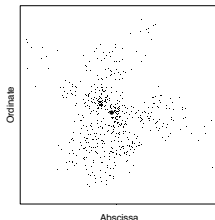
$$R^{max} = \max\{R_{jn} : j \in \mathcal{J}, n \in \mathcal{N}_j\}$$

where $R_{jn} = \|X_{jn}\|$

$\mathcal{J} = \{1, \dots, J\}$ if $J > 0$ and the empty set otherwise

$\mathcal{N}_j = \{1, \dots, N_j\}$ if $N_j > 0$ and the empty set otherwise

By convention, if no particle is released ($J = 0$ or $N_j = 0$ for all j), then $R^{max} = 0$



$$R^{max} = \max\{R_{jn} : j \in \mathcal{J}, n \in \mathcal{N}_j\}$$

Under the GDM and IDMs, the distribution of the distance between the origin and the furthest deposited propagule is zero-inflated and satisfies:

$$P(R^{max} = 0) = \exp[\lambda\{p_{\mu, \sigma^2}(0) - 1\}]$$
$$f_{R^{max}}(r) = \lambda f_{R_j^{max}}(r) \exp\{\lambda(F_{R_j^{max}}(r) - 1)\}, \quad \forall r > 0,$$

where $f_{R_j^{max}}$ is the p.d.f. of the distance $R_j^{max} = \max\{R_{jn} : n \in \mathcal{N}_j\}$ between the origin and the furthest deposited propagule of group j , and $F_{R_j^{max}}$ is the corresponding cumulative distribution function ($F_{R_j^{max}}(r) = P(R_j^{max} = 0) + \int_0^r f_{R_j^{max}}(u) du$).

→ Distribution of R_j^{max} ?

Under the IDMs, $N_j = 1$ for all $j \in \mathcal{J}$ and, consequently,
 $p_{\mu, \sigma^2}(0) = 0$ and

$$\begin{aligned} f_{R_j^{\max}}(r) &= f_{R_{j_n}}(r) \\ &= \begin{cases} \int_0^{2\pi} r f_{X_{j_n}}((r \cos \theta, r \sin \theta)) d\theta & \text{for the IDM1} \\ \int_0^{2\pi} r f_{X_j}((r \cos \theta, r \sin \theta)) d\theta & \text{for the IDM2.} \end{cases} \end{aligned}$$

Under the GDM, the distribution of R_j^{max} is zero-inflated and satisfies:

$$P(R_j^{max} = 0) = p_{\mu, \sigma^2}(0)$$

$$\begin{aligned} f_{R_j^{max}}(r) &= \int_{\mathbb{R}^2} f_{R_j^{max}|X_j}(r | x) f_{X_j}(x) dx \\ &= \sum_{q=1}^{+\infty} q p_{\mu, \sigma^2}(q) \int_{\mathbb{R}^2} f_{R_{jn}|X_j}(r | x) F_{R_{jn}|X_j}(r | x)^{q-1} f_{X_j}(x) dx, \end{aligned}$$

where $f_{R_{jn}|X_j}$ is the conditional distribution of R_{jn} given X_j satisfying

$$f_{R_{jn}|X_j}(r | x) = 2r \int_0^{r^2} h_1(u, x) h_2(r^2 - u, x) du,$$

$$h_i(u, x) = \frac{f_i(\sqrt{u}, x) + f_i(-\sqrt{u}, x)}{2\sqrt{u}}, \quad \forall i \in \{1, 2\},$$

$$f_i(v, x) = \frac{1}{\sqrt{2\pi\nu\|x\|}} \exp\left(-\frac{(v - x^{(i)})^2}{2\nu\|x\|}\right), \quad \forall i \in \{1, 2\},$$

$x = (x^{(1)}, x^{(2)})$ and $F_{R_{jn}|X_j}(r | x) = \int_0^r f_{R_{jn}|X_j}(s | x) ds$.

Theorem

Consider a GDM and an IDM1 characterized by the same parameter values except that $E(J) = \tilde{\lambda}$, $E(N_j) = \tilde{\mu}$ and $V(N_j) = \sigma^2$ for the GDM, and $E(J) = \tilde{\lambda}\tilde{\mu}$, $E(N_j) = 1$ and $V(N_j) = 0$ for the IDM1 (\Rightarrow same marginal dispersal kernel). Then, for all $r > 0$ the probability $P(R^{\max} \geq r)$ is lower for the GDM than for the IDM1.

Theorem

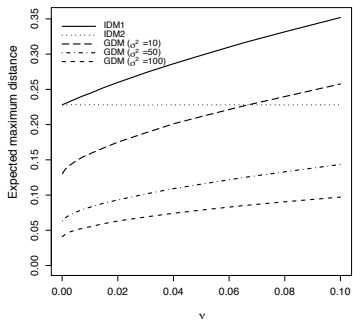
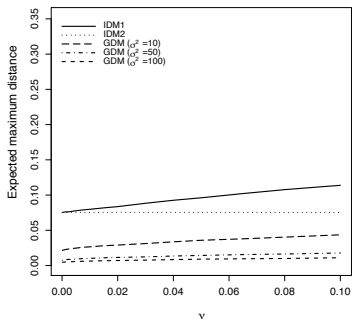
Consider an IDM1 and an IDM2 characterized by the same parameter values except that $\nu > 0$ for the IDM1 and $\nu = 0$ for the IDM2.

Then, for all $r > 0$ the probability $P(R^{\max} \geq r)$ is lower for the IDM2 than for the IDM1.

Interpretation:

The population of particles are less concentrated in probability for the IDM1 than for the GDM and the IDM2

$E(R^{max})$:

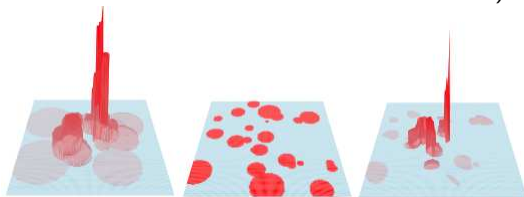


Conclusion

- ▶ With group dispersal, one can generate multiple foci whereas the particles are more concentrated

Perspectives

- ▶ Toward analytic results about the farthest particle in the spatio-temporal case
(→ speed of propagation of epidemics)
- ▶ Inference (with Tomas Mrkvicka and Eyoub Sidi)
- ▶ Alternative representations of group dispersal (Cylinder-based models, with Tomas Mrkvicka and Antti Penttinen)



- ▶ Study of the evolutionary dynamics between group dispersal and independent dispersal